Linear Regression II

Introduction to Statistics

I CAN'T BELIEVE SCHOOLS ARE STILL TEACHING KIDS ABOUT THE NULL HYPOTHESIS. I REMEMBER READING A BIG STUDY THAT CONCLUSIVELY

DISPROVED IT HEARS AGO.



* Recap of Multiple Linear Regression.

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- * Inference in Linear Regression.
 - * Estimating the *uncertainty* of OLS estimates: standard error of the regression coefficient and confidence intervals.
 - * Testing hypotheses in OLS: *t*-statistics and *p*-values.
 - * As last time: build up from intuitions about simplest cases.

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 - * As last time: build up from intuitions about simplest cases.
- * Finishing up with OLS Assumptions: two more conditions for inference with OLS.

Regression: Recap

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- * Where each β_j represents the average increase in *Y* associated with a one-unit increase in X_j holding the other variables constant.
- * How do we pick the coefficients?
- * The most common method (not the only one!) is Ordinary Least
 Squares (OLS) choose the combination of coefficients that
 minimise the sum of squared residuals.

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- * Each observation *i* will have its own residual $\hat{\epsilon}_i = Y_i \hat{Y}_i$

* So OLS will choose $Y = \hat{\alpha} + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \hat{\beta}_3 X_3 \dots \hat{\beta}_p X_p + \hat{\epsilon}$ so that $\sum_{i=1}^n \hat{\epsilon}_i^2 = \sum_{i=1}^n (Y - \hat{Y}_i)^2$ is minimised.

	Dependent variable:
	Life Satisfaction (0–10)
Age	0.013*** (0.004)
Income Decile	0.163*** (0.019)
Female	0.288*** (0.100)
Religiosity (0–10)	0.022 (0.017)
Years of Education	-0.003 (0.014)
Divorced	-0.354 (0.299)
Single	-0.118 (0.131)
Widowed	-0.412** (0.189)
Constant	5.713*** (0.321)
Observations	1,601
R ²	0.078
Adjusted R ²	0.073
Residual Std. Error	1.947 (df = 1592)
F Statistic	16.778^{***} (df = 8; 1592)

*p<0.1; **p<0.05; ***p<0.01

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- * R²: the model explains (R²) × 100 % of the variance in *Y*.

Inference in Linear Regression

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- * We observe a **sample mean**. How does it relate to the **population mean**?
- * Measures of uncertainty:
 - Standard Error of the sample mean: estimated std.
 deviation of the sample mean across repeated sampling from the population.
 - * 95% Confidence Interval: range of values which in
 95% of the samples includes the population mean.

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- * Linear model: theory of the data-generating process in the **population** (informally: in the 'real world')
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- * Linear model: theory of the data-generating process in the **population** (informally: in the 'real world')
 - * We assume *Y* is a linear function of *X*s (systematic component) plus chance *ε* (random/stochastic component).
- * OLS is an **estimator**: produces estimates from the data (the **sample**) of unobserved **population** parameters.
 - * Just like sample means. But in OLS we get more than one estimate of more than one parameter: $\hat{\alpha}, \hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3...$

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- * This sampling framework allows us to (1) quantify the *uncertainty* of our OLS estimates, and (2) test the *statistical significance* of the relationships they express.




	Intercept	Slope
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Back to Pct. Leave = $\alpha + \beta$ Pct. Degree + ϵ

Over Many Repeated Samples... Intercept Slope **Population** -1.05 81.39 Sample 1 -1.10 81.69 60 Sample 2 77.51 -0.93 Percent Leave Sample 3 79.12 -0.97 Sample 4 83.58 -1.15 40 Mean of Sample **~ 81.39 →** -1.05 **Estimates** Std. Dev of 20 Sample $SE(\alpha)$ **SE(β)** 0 10 20 30 40 50 Percent Degree

Estimates

Std. Errors of OLS Coefficients in R

* With the summary () function:

```
model1 <- lm(data = brexit, percent leave ~ percent degree)</pre>
summary(model1)
##
## Call:
## lm(formula = percent leave ~ percent degree, data = brexit)
##
## Residuals:
## Min 1Q Median 3Q Max
## -23.855 -2.462 2.203 4.819 11.175
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 81.6906 2.8560 28.60 <2e-16 ***
## percent degree -1.0982 0.1063 -10.33 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

Std. Errors of OLS Coefficients in R

```
* Tidied up with stargazer()
```

```
stargazer(model1, type = "text")
##
##
##
                       Dependent variable:
##
                       percent_leave
##
##
                       -1.098***
## percent degree
##
                            (0.106)
##
                           81.691***
## Constant
                            (2.856)
##
##
##
## Observations
                             100
## R2
                             0.521
## Adjusted R2
                             0.517
## Residual Std. Error 7.099 (df = 98)
## F Statistic 106.771^{***} (df = 1; 98)
##
```

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- * It expresses the **uncertainty** of the estimated coefficient.
- The problem: we do not observe the population. But we can estimate it from the sample by making some assumptions about the nature of the error term (*c* without a hat) in the population.

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$$\hat{\sigma}^2 = \operatorname{Var}(\hat{\epsilon}^2) = \frac{\sum (\hat{\epsilon}_i - \operatorname{mean}(\hat{\epsilon}))^2}{n-2} = \frac{\sum \hat{\epsilon}^2}{n-2}$$

S.E.
$$(\hat{\beta}) = \sqrt{\frac{\sum \hat{\epsilon}^2}{(\frac{1}{n-2})\sum (x_i - \bar{x})^2}}$$

* Substituting in σ^2 ...

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- * Your standard errors will be larger if...
 - * X does a poor job at predicting Y ($\sum e^2$ goes up)
 - * X does not vary much ($\sum (x_i \bar{x})^2$ goes down)
 - * Your sample is small ($\frac{1}{n-2}$ goes down).

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* Each variable $(X_1, X_2...)$ will have an associated coefficient $(\hat{\beta}_1, \hat{\beta}_2)$, which in turn come with their own standard error.



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- * Each variable $(X_1, X_2...)$ will have an associated coefficient $(\hat{\beta}_1, \hat{\beta}_2)$, which in turn come with their own standard error.
- * Std. Error of $\hat{\beta}_2 \rightarrow$ estimated std. deviation of the slope of X_2 across repeated hypothetical sampling.


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- ★ Different critical values → difference confidence levels.
- * The confidence interval thus calculated will include the 'true' population slope β in 95% of the (hypothetical, random) samples.

```
model1
##
## Call:
##
  lm(formula = percent leave ~ percent degree, data = brexit)
##
## Coefficients:
##
  (Intercept) percent degree
                          -1.098
##
          81.691
confint(model1)
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* We also get confidence intervals for $\hat{\alpha}$ — but we're normally interested in the uncertainty of our slope.





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Population	-1.05			
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Over many repeated samples	→ -1.05			In 95% (19 out of 20) samples

	Dependent variable:
	Life Satisfaction (0–10)
Age	0.013*** (0.004)
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- * Drawback: predictors may be on very different scales.
- Makes most sense when you have all categorical predictors (e.g. conjoint experiment).





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- * Drawback: categorical variables make little sense.



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14 years of education, religiosity = 3

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t-statistic of a sample mean =
$$\frac{\bar{X} - X_0}{SE_X}$$

p-value (two-tailed): probability of obtaining a test statistic at least as extreme as the one we observe, under the null hypothesis.
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- * But if the slope is estimated **from a sample**, how sure can we be that the relationship it expresses is really there **in the population**? With a *t*-test!
- * Maths to make this work require an additional assumption: that the **error term is normally distributed**, i.e. $\epsilon \sim \mathcal{N}(0, \sigma^2)$.

##	Coefficients	5:								
##		Estimate	Std. Error	t value	Pr(> t)					
##	(Intercept)	6.5526	0.2173	30.150	<2e-16	***				
##	religiosity	0.1053	0.0471	2.236	0.0262	*				
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- * From a sample of size *n*, we estimate $\hat{\beta} = 0.1053$.
- * Assume a population where *X* and *Y* are completely uncorrelated, $Y_i = \alpha + 0X + \epsilon_i$ where $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$.

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##	Signif. code	es: 0 '**	<pre> *' 0.001 ';</pre>	**' 0.01	'*' 0.05	'.'	0.1	T	T	1

- * From a sample of size *n*, we estimate $\hat{\beta} = 0.1053$.
- * Assume a population where *X* and *Y* are completely uncorrelated, $Y_i = \alpha + 0X + \epsilon_i$ where $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$.
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##	Coefficients	S:								
##		Estimate	Std. Error	t value	Pr(> t)					
##	(Intercept)	6.5526	0.2173	30.150	<2e-16	***				
##	religiosity	0.1053	0.0471	2.236	0.0262	*				
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- * We don't know the 'true' σ^2 , so we approximate it from the observed variance of the residuals $\hat{\sigma}^2$.
- * If the 'true' $\beta = 0$, how likely is it that, over many samples of size *n*, we get a slope as extreme as $\hat{\beta}$? (i.e. $\hat{\beta}_s > 0.1053$ or $\hat{\beta}_s < -0.1053$)

Life Satisfaction = $\alpha + \beta$ Religiosity + ϵ

Slope

Our Data

0.105

Life Satisfaction = $\alpha + \beta$ Religiosity + ϵ



	Slope
Our Data	0.105
Population under the null	0

Clara

Life Satisfaction = $\alpha + \beta$ Religiosity + ϵ



	olope
Our Data	0.105
Population under the null	0

Slone



	Slope
Our Data	0.105
Population under the null	0
Sample 1 from pop.	0.019

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Slopo



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Our Data	0.105
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Mean Over many repeated samples	→ 0

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Our Data	0.105
Population under the null	0
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Sample 3 from pop.	-0.011
Mean Over many repeated samples	→ 0
Std. deviation of estimates over many repeated samples	≈ SE(β)











Sample t-values from Population under Null



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- * Conventionally, 95% or $\alpha = 0.05$.
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- 4. What's the critical value?
- * For $\alpha = 0.05$, this will be about 1.96 (a bit higher when we have small samples or many predictors).
- ∗ Under the null, in 5% of the samples we will get tstatistics over 1.96 or below −1.96.

5. Is the absolute value of *t*-statistic larger or equal than the critical value?

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- * Why the absolute value? Because when the estimate is negative, the *t*-statistic will also have a negative sign.

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t-statistic and *p*-value in R

```
summary(model1)
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
                            2.8560 28.60 <2e-16 ***
## (Intercept) 81.6906
## percent degree -1.0982 0.1063 -10.33 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Coefficients:
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## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

p-values in Regression Tables

	Dependent variable:
	Life Satisfaction (0–10)
Age	0.013*** (0.004)
Income Decile	0.163*** (0.019)
Female	0.288*** (0.100)
Religiosity (0–10)	0.022 (0.017)
Years of Education	-0.003 (0.014)
Divorced	-0.354 (0.299)
Single	-0.118 (0.131)
Widowed	-0.412** (0.189)
Constant	5.713*** (0.321)
Observations	1,601
R ²	0.078
Adjusted R ²	0.073
Residual Std. Error	1.947 (df = 1592)
F Statistic	16.778*** (df = 8; 1592)

*p<0.1; **p<0.05; ***p<0.01

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* "The relationship is significant at the 99.99% level, so it's likely true / causal / worth caring about."

Don't Be This Guy

* "In 2020, Biden's tabulated votes (2,474,507) were much greater than Clinton's in 2016. [...] I tested the hypothesis that the performance of the two Democrat [sic] candidates were statistically similar by comparing Clinton to Biden. [...] I use the calculated Z-score to determine the p-value [...]. This value corresponds to a confidence that I can reject the hypothesis many times more than one in a quadrillion times that the two outcomes were similar."

(Charles Cicchetti, Lawsuit filed by the State of Texas)

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- 4. Cutoffs are arbitrary (and bad for science): p = 0.049 is just as good as p = 0.051. **Don't p-hack your way to significance**.
- 5. **Non-significant findings are valuable**. Especially if we can be very confident about the fact that there's probably no meaningful relationship ('precise null').

Least Squares Assumptions: An Essential Checklist

1. Linearity

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 - * The model **in the population** (the 'true' model) can be written as a linear combination of variables and coefficients: $Y = \alpha + \beta_1 X_1 + \beta_2 X_2 \dots \beta_p X_p + \epsilon$.

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If assumptions 1–4 are satisfied, our OLS coefficient estimates are unbiased

Classical Linear Model Assumptions

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- 3. No Perfect Collinearity
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- * We may diagnose that this assumption is violated (heteroskedasticity) from plotting the residuals against the independent variables.
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- * One popular fix: heteroskedasticity-consistent standard errors (more conservative than default standard errors).

Violation of Homoskedasticity Assumption



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 - * In large samples, we can invoke the central limit theorem to conclude that the error term approximates a normal distribution.
 But no easy fix in small samples.
 - * Non-normal errors are usually the result of linearity assumption not holding (e.g. Y can only take a limited number of values). If you fix that, things are usually fine.



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- * These especially **homoskedasticity** rarely hold in observational studies, so 'default' S.E. and *p*-values are likely wrong (usually, too small).
- * Next week: moving beyond linear additive relationships

Thank you for your kind attention!

Leonardo Carella leonardo.carella@nuffield.ox.ac.uk